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## Abstract

27 Population responses to environmental variation ultimately depend on withinindividual and among-individual variation in labile phenotypic traits that affect fitness, 28 29 and resulting episodes of selection. Yet, complex patterns of individual phenotypic 30 variation arising within and between time periods, and associated variation in selection, 31 have not been fully conceptualised or quantified. We highlight how structured patterns 32 of phenotypic variation in dichotomous threshold traits can theoretically arise and 33 experience varying forms of selection, shaping overall phenotypic dynamics. We then fit novel multistate models to ten years of band-resighting data from European shags to 34 35 quantify phenotypic variation and selection in a key threshold trait underlying spatioseasonal population dynamics: seasonal migration versus residence. First, we 36 37 demonstrate substantial among-individual variation alongside substantial between-year individual repeatability in within-year phenotypic variation ('flexibility'), with weak 38 39 sexual dimorphism. Second, we demonstrate that between-year individual variation in within-year phenotypes ('supraflexibility') is structured and directional, consistent with 40 41 the threshold trait model. Third, we demonstrate strong survival selection on within-42 year phenotypes, and hence on flexibility, that varies across years and sexes, including 43 episodes of disruptive selection representing costs of flexibility. By quantitatively combining these results, we show how supraflexibility and survival selection on 44 migratory flexibility jointly shape population-wide phenotypic dynamics of seasonal 45 46 movement.

#### Introduction

48 Population responses to environmental variation and change ultimately arise through 49 within-individual and among-individual variation in phenotypic expression of key 50 environmentally-sensitive traits, and resulting variation in fitness components and consequent selection (Nussey et al. 2007; Gienapp et al. 2008; Chevin et al. 2010; Fox 51 52 et al. 2019). Quantifying temporal dynamics of individual phenotypes within and 53 between consecutive time periods, and quantifying associated dynamics of selection, is 54 consequently necessary to understand and predict joint phenotypic, population dynamic 55 and micro-evolutionary outcomes across temporal scales (Hendry 2016; Araya-Ajoy et 56 al. 2015; Arnold et al. 2019; de Villemereuil et al. 2020).

57 Rapid environmentally-induced changes in phenotypes occurring on short 58 (within-generation) timescales typically primarily reflect within-individual responses, 59 representing phenotypic plasticity of labile traits (Sih et al. 2011; Gabriel et al. 2005; 60 Charmantier et al. 2008; Chevin et al. 2010). In wild settings where all pertinent axes 61 of underlying environmental variation and resulting multi-dimensional reaction norms are hard to explicitly quantify, such plastic responses can be usefully summarised as 62 63 patterns and magnitudes of observed within-individual phenotypic variation (here 64 termed 'phenotypic flexibility' or simply 'flexibility'; Piersma and Drent 2003; Glossary and summary of key terms in Online Supplemental Material, 'OSM', S1). 65 66 Hierarchical levels of flexibility could then arise across timescales; individuals could be phenotypically invariable or variable within time periods (e.g. within years), while 67 68 their within-period sequence of phenotypes could vary between periods (e.g. between 69 years). Such expression of different levels of phenotypic flexibility arising across the 70 hierarchy of timescales could then experience episodes of selection, which could be 71 consistent or divergent, and could differ between groups of individuals (e.g. sexes),

72 and/or vary in relation to environmental conditions. However, despite their central roles 73 in shaping population responses, key aspects of expected and realised hierarchical 74 patterns of phenotypic flexibility across nested temporal scales, and associated patterns 75 of selection, have not yet been fully conceptualised or quantified for key traits that directly link phenotypic and population dynamic outcomes. Here, we provide a 76 conceptual framework to rationalise hierarchical patterns of phenotypic flexibility and 77 78 selection on flexibility concerning dichotomous threshold traits. We apply this 79 framework to estimate key effects for the ecologically critical threshold trait of seasonal 80 migration versus residence.

81 Recent studies focussing primarily on traits that are continuously distributed on observed phenotypic scales have established that reaction norm slopes defining 82 83 phenotypic plasticity can vary among individuals within populations (Nussey et al. 84 2007; Brommer 2013; Dingemanse and Wolf 2013; Hendry 2016). Some individuals consequently show more phenotypic variation than others across a given time period or 85 86 axis of environmental variation. Further, such among-individual differences may be repeatable (Araya-Ajoy and Dingemanse 2017; Mitchell and Biro 2017; Cornwell et al. 87 88 2019; Strickland et al. 2021), meaning that individuals show consistent degrees of 89 relative phenotypic variation. Nonetheless, reported repeatabilities of individual 90 reaction norm slopes of ~0.4-0.5 (Araya-Ajoy and Dingemanse 2017; Mitchell and Biro 2017; Strickland et al. 2021) also imply potential for substantial within-individual 91 92 variation in phenotypic variation. Individuals could consequently vary in the extent to 93 which they change their phenotype between time periods or environmental axes. Yet, 94 key aspects of the potential for structured forms of among-individual variation in 95 within- versus between-period phenotypic variation, and associated forms of selection, still remain to be conceptualised and quantified. 96

97 In particular, numerous labile behavioural and life-history traits that link environmental changes with population outcomes are expressed as discrete alternative 98 99 phenotypes rather than as continuously distributed phenotypes (e.g. dominant versus 100 subordinate behaviour, breeding versus skipping, mate fidelity versus divorce, or 101 seasonal migration versus residence; Snell-Rood et al. 2018; Reid and Acker 2022). 102 Such traits can often be appropriately conceptualised as quantitative genetic 'threshold 103 traits'. Here, an underlying continuously distributed latent variable termed 'liability', 104 which can comprise combinations of genetic and environmental effects, translates into 105 expression of discrete alternative phenotypes X or Y when below versus above some 106 threshold value (fig. 1A; Gianola 1982; Roff 1996; Lynch and Walsh 1998 Ch.25). 107 Variation occurring on the phenotypic scale can then be decoupled from reaction norm 108 slopes defining liability-scale plasticity, and instead reflect variation in liability-scale 109 reaction norm elevations. There can consequently be among-individual variation in observed phenotypic flexibility without any among-individual variation in liability-scale 110 111 reaction norm slopes, and conversely there can be among-individual variation in liability-112 scale reaction norm slopes that does not result in among-individual variation in phenotypic 113 flexibility (Reid and Acker 2022). Such threshold traits could consequently generate distinctive and biologically important patterns of phenotypic variation within and 114 115 among individuals, within and among temporal scales, that could in turn experience 116 varying selection (fig. 1, 2).

117 Specifically, within any focal time period over which selection could act, 118 individuals might inflexibly express phenotype X or phenotype Y, or flexibly switch 119 between the two (fig. 1A,B). The set of possible within-period phenotypic sequences 120 (hereafter 'tactics') constitutes a continuum of within-individual expression of X versus 121 Y (summarized as 'full-X', 'mixed-XY' and 'full-Y'), reflecting underlying individual

122 liability elevations (fig. 1B,C, 2A). Then, between any two consecutive time periods, individuals could either retain the same tactic, or switch tactic, and thereby potentially 123 alter their degree of flexibility (fig. 2A). We term such between-period variation 124 125 'supraflexibility' to stress the hierarchy of phenotypic expression, and hence of flexibility (OSM S1). Such hierarchical timeframes are necessarily relative, meaning 126 that the relevant duration of periods through which flexibility and supraflexibility are 127 128 quantified will be system-, trait- and question-specific. Given the threshold trait model, 129 the distance of an individual's initial liability from the threshold (i.e. the elevation) will 130 affect whether or not subsequent environmental variation will cause its liability to cross 131 the threshold, and thereby cause phenotypic change (fig. 1B; Reid and Acker 2022). 132 Accordingly, flexibility and supraflexibility are likely to be structured and directional, 133 such that some transitions occur more frequently than others. Specifically, full-X and 134 full-Y individuals should be more likely to switch to mixed-XY than to full-Y and full-X respectively. Supraflexibility should also intrinsically covary with flexibility, such 135 136 that individuals that are more flexible within periods (i.e. mixed-XY) could also on 137 average be more likely to subsequently switch to full-X or full-Y than vice versa (fig. 138 2A). Flexibility and supraflexibility in phenotypic expression of a labile threshold trait can therefore represent different manifestations of the same underlying reaction norm 139 140 enacted on the latent liability scale, but occurring on hierarchical timeframes (e.g. Reid 141 and Acker 2022).

Selection could then act on flexibility and/or on supraflexibility if individual tactics affect relative survival and/or reproduction (fig. 2B,C). For example, inflexible individuals that express full-X or full-Y during a given period could potentially have similar survival probabilities, while flexible individuals that express mixed-XY could have higher or lower survival probability. These scenarios could represent benefits or

147 costs of phenotypic flexibility, and generate stabilizing or disruptive selection along the X-Y continuum (fig. 2B). Yet, in other periods, environmental conditions could 148 potentially mean that survival is higher for full-Y than full-X and intermediate for 149 150 mixed-XY, generating directional selection for Y on the X-Y continuum (or, 151 conversely, directional selection for X; fig. 2B). Here, supraflexible individuals that 152 switched from full-X to mixed-XY, or from mixed-XY to full-Y, between two such 153 consecutive periods would then have higher overall survival probability than individuals that did not switch, effectively generating selection for supraflexibility (fig. 154 155 2C). Yet, such effects of sequences of selection on flexibility in generating selection on supraflexibility could be further compounded by direct selection against 156 157 supraflexibility, for example manifested as lower survival of individuals that is directly 158 caused by switching tactics between periods. The overall shape and net strength of 159 selection acting on supraflexibility could then vary across time (e.g. fig. 2B,C), driving complex short-term phenotypic dynamics within and across time periods, as well as 160 161 potentially shaping longer-term evolutionary change.

162 Specifically, episodes of survival selection could directly cause additional within-163 generation shifts in population-wide frequencies of phenotypic tactics, and hence of 164 flexible and inflexible individuals. The magnitude of such shifts will depend on amongindividual variance in relative fitness (i.e. the opportunity for selection; Arnold and 165 Wade 1984), which in turn depends on relative tactic frequencies prior to selection. 166 167 Resulting within-generation shifts could persist across time periods if tactics are highly repeatable, or could be quickly mitigated or even reoriented, depending on the 168 169 magnitude and directionality of supraflexibility that non-randomly alters tactic 170 frequencies among surviving individuals. Moreover, episodes of sex-specific selection could directly generate or accentuate sex biases in tactic frequencies, reflecting sexual 171

172 dimorphism in underlying liability and implying sexual dimorphism in flexibility 173 and/or supraflexibility. Thus, while sex-specific selection in continuously distributed phenotypic traits can generally shape evolutionary change and drive evolution of sexual 174 175 dimorphism (Lande 1980; Kruuk et al. 2008; Cox and Calsbeek 2009), such outcomes can be further affected by the properties of threshold traits (Reid and Acker 2022). 176 Specifically, if there is sexual dimorphism in phenotypic expression of the threshold 177 178 trait, there will be intrinsic sexual dimorphism in intensity of selection (Falconer and 179 Mackay 1996). Overall, therefore, understanding and predicting phenotypic and 180 population outcomes requires quantifying sex-specific frequencies of flexible and 181 inflexible tactics, and associated patterns of supraflexibility and temporal variation in 182 survival selection, regarding key environmentally-sensitive dichotomous traits that 183 drive population dynamics. But, such effects have not previously been explicitly 184 quantified, or subsequently combined to evaluate relative effects on population-level phenotypic variation. 185

186 One ecologically critical threshold trait is seasonal migration (i.e. reversible movement across seasons, hereafter 'migration'; Dingle 1996; Newton 2008) versus 187 year-round residence in populations with facultative migration. Resulting 'partial 188 189 migration', where sympatric-breeding individuals can express alternative resident and 190 migrant phenotypes (hereafter 'R' and 'M') through non-breeding seasons, occurs in 191 diverse taxa including fish, mammals, birds, reptiles, amphibians (Lundberg 1988; Chapman et al. 2011; Reid et al. 2018). Because R and M individuals are by definition 192 193 seasonally spatially segregated, local environmental variation can cause episodes of 194 strong survival selection, notably during extreme climatic events ('ECEs') that affect 195 some locations more than others (Acker et al. 2021a). Further, various annual migratory tactics may be expressed, such that some individuals inflexibly express R or M 196

197 throughout the non-breeding season ('full-R' or 'full-M'), while others flexibly switch 198 between R and M ('mixed-RM'), for example manifested as late departure from breeding areas (Cagnacci et al. 2011; Fudickar et al. 2013; Reid et al. 2020). Any 199 200 change in phenotypic expression along the R-M continuum, and hence in population-201 wide frequencies of full-R, mixed-RM, and full-M, will then directly alter spatio-202 seasonal population dynamics (Reid et al. 2018). However, even though variation in 203 expression of migration versus residence is currently attracting considerable attention 204 and has long been conceptualised as a threshold trait (Berthold 1988; Pulido et al. 1996; 205 Dodson et al. 2013), the hierarchical structure of sex-specific flexibility and 206 supraflexibility, and associated selection, have not been fully quantified.

207 To achieve these ambitions, we fitted a novel multistate capture-recapture model 208 to ten years of large-scale year-round capture-resighting data from individually marked 209 adult European shags (Gulosus aristotelis, hereafter 'shags'). Our model represents individual phenotypic variation along the R-M continuum as a dynamic mixture of 210 211 annual tactics, allowing us to quantify four sets of effects. First, we estimated year- and sex-specific tactic frequencies, and thereby quantified the degree of among-individual 212 213 variation in within-year phenotypic flexibility in migration versus residence. Second, 214 we estimated probabilities of all possible forms of tactic switching between consecutive 215 years, and hence quantified sex-specific tactic repeatability and structured supraflexibility. Third, we estimated tactic-dependent survival probabilities, hence 216 quantified temporal variation in the form of sex-specific survival selection on migratory 217 tactics, and hence on migratory flexibility. Finally, we combined all these estimates to 218 219 quantify relative effects of supraflexibility and survival selection on flexibility, 220 alongside entry of new adults into the analysed population, on population-wide changes in tactic frequencies (including flexibility) from each year to the next. We thereby 221

illustrate how phenotypic dynamics in a key trait that directly affects seasonal population distributions can be shaped by combinations of sex-specific flexibility, supraflexibility, and selection on flexibility in the context of among-year environmental variation. We discuss these outcomes in the context of general theoretical expectations for phenotypic and micro-evolutionary dynamics of plastic threshold traits.

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# Material and methods

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#### Study system and data collection

230 Quantifying variation in phenotypic flexibility and supraflexibility of migration versus 231 residence, and associated survival selection, requires repeated measurement of 232 locations and survival within and across multiple years in numerous individuals. We collected the required data using visual field resightings during 2009–2019 in partially-233 234 migratory shags that breed on Isle of May National Nature Reserve (hereafter 'IoM'), 235 Scotland (56°11'N, 2°33'W; Daunt et al. 2014; Grist et al. 2014; Keogan et al. 2021). 236 Shags are seabirds that must return to shore every day to dry their partially-wettable 237 plumage (Harris and Swann 2002), and migrants in our focal population typically winter 100-500 km from IoM. Marked individuals can consequently be resighted 238 onshore within the UK throughout each year, allowing direct observation of 239 240 individuals' winter locations (Grist et al. 2014; Acker et al. 2021a), and hence allowing inference on resident versus migrant phenotypes and resulting annual migratory tactics. 241

To generate a population of marked adults, ≥80% of chicks hatched on IoM since
1997 were marked before fledging with an individually-coded colour band (field
readable up to ~150m with a telescope) and an inscribed metal band. Since ~90% of

245 individuals that survived to adulthood recruited locally (median age 3 years; Barlow et al. 2013), by 2009 most breeding adults were banded. Remaining unbanded adults, 246 including incoming immigrants, were caught and banded during breeding as far as 247 248 feasible. Every non-breeding season (early September to late February, hereafter 'winter') from 2009-10 to 2018-19, resighting surveys were carried out approximately 249 biweekly within the residency area (IoM and adjacent day roosts) and at key alternative 250 251 wintering locations across eastern Scotland, with additional resightings collected 252 elsewhere (Grist et al. 2014; Acker et al. 2021a; summarised in OSM S2). During 253 subsequent breeding seasons (April-June 2010 to 2019), nests and roost sites on IoM 254 were surveyed at least twice weekly to identify colour-banded adults, resulting in very 255 high breeding season resighting probabilities (mean 0.95, range 0.90-0.98) facilitating 256 inferences on annual survival (Acker et al. 2021a). Sexes of most adults were 257 determined through vocalizations and/or genotyping (Acker et al. 2021a).

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## Model design

260 In general, unbiased estimation of phenotypic variation and associated selection 261 requires analyses that account for processes leading to non-random missing data due to inevitable imperfections in abilities to measure or assign individual phenotypes and/or 262 associated fitness (Gimenez et al. 2008; Hadfield 2008; Nakagawa and Freckleton 263 264 2008). Highly non-random missing data inevitably arise when the focal trait is 265 migration versus residence inferred from field observations. Here, marked individuals can be missed or move outside surveyed areas, causing resighting failure that varies 266 267 according to phenotype given spatio-temporal variation in observational efficiency. Missing data following the last resighting of an individual could also reflect mortality, 268

269 confounding estimates of phenotype-dependent resighting and survival probabilities (and hence of selection). Further, our current focal phenotypic tactics (namely season-270 271 long full-R, full-M and mixed-RM) can never be instantaneously observed, but must be 272 inferred from sequences of observations, which are often incomplete due to resighting failure and/or within-winter mortality. Consequently, (generalized) linear mixed 273 models that are widely used to estimate within- and among-individual phenotypic 274 275 variation in readily observed traits (e.g. Araya-Ajoy et al. 2015) cannot be directly applied to estimate migratory flexibility, supraflexibility and associated selection, at 276 277 least without an observation model to account for non-random missing data.

278 Hence, to answer our current questions, we devised a novel full-annual-cycle multistate model that uses capture-recapture methods to enable robust probabilistic 279 280 inference on variable phenotypic expression of residence versus migration and 281 associated survival from large-scale mark-recapture data given observation failure that is non-random with respect to both tactic and survival. Specifically, our model 282 283 represents the process of within-year transitions between states defined by current individual location, parameterised by movement and survival probabilities, and 284 285 overlaid by the observation process parameterised by resighting probabilities. By 286 including finite individual heterogeneity defined through constraints on transitions between resident and migrant states within years, we explicitly modelled within- and 287 among-individual phenotypic variation (and hence variation in flexibility and 288 supraflexibility) as a dynamic mixture of annual migratory tactics parameterised by 289 290 initial tactic probabilities and tactic switching probabilities (fig. 3).

To maximise use of available data, we divided the annual cycle into five resighting occasions comprising the breeding season (occasion 1) and four occasions through the subsequent winter (occasions 2–5, fig. 3; OSM S2). We defined states

294 representing six spatially-distinct areas: the 'residency area' (IoM area), four observed migrant areas (defined based on geography and survey frequency), and one unobserved 295 migrant area comprising unsurveyed migrant destinations (i.e. a 'ghost area'; OSM S2). 296 297 This spatial structure allowed us to account for spatio-temporal variation in resighting probability and hence minimise related bias (Acker et al. 2021a). Accordingly, alive 298 299 individuals can be resighted where they are located, or not resighted, according to sex×occasion×year×area-dependent ('×' denotes interacting effects) resighting 300 301 probability (p, which is zero in the ghost area; OSM S3). This requires that migrants 302 are spatially distributed according to probabilities of moving to a specific migratory 303 area conditional on departure ( $\delta$ , occasion×year×area-dependent) and of moving 304 between migratory areas conditional on not returning to the residency area ( $\sigma$ , assumed 305 constant across occasions and years; Acker et al. 2021a; OSM S3). This spatial division 306 of migrant destinations does not affect the current focus on dichotomous migrant versus 307 resident phenotypes (Chapman et al. 2011; Pulido 2011; Dodson et al. 2013), defined 308 based on departure from and return to the residency area (Acker et al. 2021a).

309 Specifically, every breeding season (occasion 1, when all individuals are located 310 in the residency area), new individuals enter the dataset as adults and are envisaged to 311 undertake a full-R, mixed-RM or full-M tactic through the subsequent winter (i.e. occasions 2-5; fig. 3), according to sex×year-dependent initial tactic probabilities ( $\pi$ ; 312 OSM S3). Each tactic is characterised by specific constraints on departure from and 313 314 return to the residency area between winter occasions. Full-R individuals remain in the residency area through all annual occasions. Full-M individuals depart from the 315 316 residency area between occasion 1 and 2 and do not return until subsequent occasion 1 (fig. 3). Meanwhile, mixed-RM individuals enact a phenotypic sequence that differs 317 from both full-R and full-M (for example, by departing after occasion 2 or 3, 318

319 representing within-year phenotypic flexibility; fig. 3, OSM S3). To account for potential variation in migration timings, phenotypic variation occurs in mixed-RM 320 321 following probabilities of departing from ( $\varepsilon$ ) and returning to ( $\tau$ ) the residency area, 322 which are sex×occasion×year×area-dependent. Previous analyses of two years of data from a different partially migratory shag population showed that division into the three 323 full-R, full-M and mixed-RM migratory tactics provides a good representation of 324 325 overall phenotypic variation (Reid et al. 2020; OSM S2). Following these previous analyses and inspection of our current dataset, occasion time windows (particularly 326 327 occasion 2) were defined to appropriately distinguish the three tactics (OSM S2).

328 Between consecutive years, individuals either repeat the same tactic or switch to 329 another tactic, following sex×year×tactic-dependent switching probabilities ( $\kappa$ ). The tactic mixture is consequently dynamic, and allows for structured supraflexibility of 330 331 migration versus residence (fig. 3, OSM S3). Mortality occurs between occasions according to sex×tactic×occasion×year-dependent survival probabilities ( $\phi$ ), that are 332 333 multiplied across occasions within years to yield sex×tactic×year-dependent annual survival probabilities ( $\phi$ ; OSM S3). Since there is virtually no breeding dispersal from 334 335 IoM (Aebischer 1995; Barlow et al. 2013), these survival probabilities represent true 336 survival probabilities for focal adults (Acker et al. 2021a).

Overall, this model fulfils our current objectives of estimating sex-specific initial tactic probabilities and tactic switching probabilities and thereby estimating variation in flexibility and supraflexibility of migration versus residence (fig. 2A) in each year and sex, and of estimating associated survival probabilities and hence survival selection on flexibility (fig. 2B). Our current model does not estimate survival selection on supraflexibility (fig. 2C), which requires major future developments and data (see *Discussion*). Since previous and exploratory analyses did not reveal major age effects on tactic frequencies (Reid et al. 2020) or survival probabilities (Acker et al. 2021a) in
adults, and incorporating age effects within our model would entail substantial extra
complexity, such effects were not further considered in current analyses.

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#### Dataset and model analysis

We fitted the model to capture-resignting histories of 2304 adult shags (1185 females and 1119 males), compiled from 48,413 year-round field sightings (OSM S2). Each history comprises a sequence of observation events indicating whether and where the individual was resignted in each occasion, starting from their first breeding attempt on IoM during 2009–2018 and ending in breeding season 2019.

Previous cross-sectional analyses spanning 2009–2018 showed that individuals 354 355 that were resident versus migrant at any occasion (i.e. not accounting for annual tactics) 356 typically had similarly high survival probabilities. However, ECEs comprising extreme late-winter storms caused substantial mortality in winters 2012–13, 2013–14 and 2017– 357 18, and also caused episodes of strong survival selection against current residents in 358 359 winters 2012-13 and 2017-18 (Acker et al. 2021a). These ECEs now provide opportunities to assess whether frequencies of annual migratory tactics (and hence 360 flexibility), between-year switching (i.e. supraflexibility) and/or selection varied with 361 362 known major environmental perturbations. Accordingly, in our current analyses,  $\Phi$  for each tactic was considered to be constant across all non-ECE years, but allowed to differ 363 364 in each ECE year (2012-13, 2013-14, 2017-18). Direct ECE effects on expression of migratory tactics in 2017-18 cannot be detected by our current model because the ECE 365 366 occurred in March (i.e. after the last specified winter occasion 5), but tactic-specific 367 survival probabilities can still be estimated.

368 To estimate model parameters, we used Stan, a probabilistic programming language for Bayesian inference, called from R through package rstan (Carpenter et al. 369 2017; R core team 2020). For all parameters, we used priors that were uniform over the 370 371 natural range of probabilities or probability simplexes (OSM S3). Posterior predictive checks (Gelman et al. 1996) devised for full-annual-cycle multistate capture-recapture 372 analyses (Acker et al. 2021a) indicated good overall model fit to the data regarding 373 374 current objectives (OSM S4). Details of posterior sampling procedures and diagnostics, full posterior samples and numerical summaries, and full model code, are in OSM S5. 375

376 Sex-, year- and tactic-specific switching and survival probabilities are directly 377 estimated by the model. However, the relative population frequencies (i.e. proportions) of each migratory tactic in each year, and hence the degree of among-individual 378 379 variation in flexibility, are not directly estimated. We therefore calculated the posterior 380 distributions of tactic frequencies at the start of each year y once new individuals have entered the dataset  $(F_{1,y})$ . Since  $F_{1,y}$  jointly results from tactics initially undertaken by 381 382 new individuals and from tactic switching of surviving individuals, it can be derived from the numbers of individuals that entered the dataset up to and including the current 383 384 year and corresponding initial tactic, survival and tactic switching probabilities (OSM 385 S3). Moreover, to synthesise how overall phenotypic changes from year y to y+1 result from sequential within-year processes, we also calculated tactic frequencies 386 immediately after annual survival selection  $(F_{2,y})$ , and after subsequent tactic switching 387 388 ( $F_{3,y}$ ; fig. 7A; OSM S3). We then derived the net effects on relative tactic frequencies of survival selection  $(E_{1,y}=F_{2,y}-F_{1,y})$ , supraflexibility  $(E_{2,y}=F_{3,y}-F_{2,y})$ , and entry of new 389 individuals into the adult population ( $E_{3,y}=F_{I,y+1}-F_{3,y}$ ; fig. 7A). 390

391 To further summarize year-dependent parameters, we also derived grand means 392 and standard deviations across years (OSM S3). Estimates from the first study year (2009-10) are not shown, and were not included in cross-year summaries, because low
resighting probabilities prevented reliable inference on individual tactics (OSM S2).
However, 2009-10 was retained in the dataset to initialise the model and differentiate
subsequent entry of new individuals from individuals already present in the data (OSM
S2). We present all estimates of directly estimated and derived parameters as posterior
means with 95% credible intervals ('95%CI').

399 To explicitly test for between-tactic, between-sex and between-year differences in tactic frequencies, switching probabilities and survival probabilities, we computed 400 the posterior distributions of such differences ( $\Delta$ ). We assessed support for the sign of 401 402 each difference through the posterior probability that it was positive ( $Pr(\Delta > 0)$ ; values 403 close to 1 or 0 indicate strong evidence for positive or negative differences respectively, 404 whilst values close to 0.5 indicate no clear evidence for either). Finally, for each year 405 and sex, we assessed support for the four possible shapes of survival selection on the annual migratory tactic along the R-M continuum (fig. 2B). More precisely, we 406 407 calculated the posterior probability that selection stabilising was  $(\Pr(\Phi_{full-R} < \Phi_{mixed-RM} > \Phi_{full-M}))$ , disruptive  $(\Pr(\Phi_{full-R} > \Phi_{mixed-RM} < \Phi_{full-M}))$ , directional 408 409 towards full-R (Pr( $\Phi_{full-R} > \Phi_{mixed-RM} > \Phi_{full-M}$ )), or directional towards full-M 410  $(\Pr(\Phi_{full-R} < \Phi_{mixed-RM} < \Phi_{full-M}))$ . These four probabilities sum to 1, and particular probabilities close to 1 indicate strong evidence for the corresponding shape of 411 selection. 412

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#### Results

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Tactic frequencies and flexibility within years

416 All three defined migratory tactics, and hence flexible and inflexible phenotypes, 417 occurred in substantial proportions among individuals alive at the start of each focal biological year (fig. 4). Full-R was typically the most frequent tactic in both sexes 418 419 (grand mean proportions across years with 95%CI: 0.42 [0.38,0.46] in females, 0.43 [0.39,0.47] in males; fig. 4; see OSM S5 for details). Meanwhile, mixed-RM was 420 421 slightly less frequent than full-M in females (grand mean proportions: 0.26 [0.23,0.30] 422 and 0.31 [0.28,0.34], respectively), but more frequent in males (0.32 [0.28,0.36] and 423 0.25 [0.22,0.28], respectively; fig. 4; OSM S5).

There was also substantial among-year variation in tactic frequencies, particularly 424 425 for full-R (fig. 4; OSM S5). Interestingly, the two years where phenotypic effects of ECEs could potentially be detected (2012-13 and 2013-14) showed relatively low 426 427 frequencies of full-R, with no evidence that full-R was more frequent than the other 428 two tactics (fig. 4; OSM S5). Together, such within-year and among-year variation indicate substantial opportunity for selection and switching among flexible and 429 430 inflexible tactics to drive appreciable shifts in phenotype frequencies, including effects 431 of ECEs.

While there was no clear evidence that the frequency of full-R differed between males and females in any year, there was strong evidence for lower frequency of mixed-RM and/or higher frequency of full-M in females than males in several years, and hence on average across years (fig. 4; OSM S5). This indicates slight sexual dimorphism in migratory tactic, implying slight dimorphism in within-winter phenotypic flexibility.

439 Probabilities of switching tactics between consecutive years revealed structured patterns of supraflexibility. Overall, probabilities of repeating the same tactic (i.e. not 440 441 switching to another tactic) were notably high (fig. 5). These probabilities were highest for full-R (grand mean: 0.86 [0.81,0.90] and 0.87 [0.83,0.90] in females and males 442 respectively), lowest for mixed-RM (grand mean: 0.54 [0.46,0.62] and 0.66 443 444 [0.60,0.73]), and intermediate for full-M (grand mean: 0.72 [0.68,0.76] and 0.74 [0.69,0.78]). There was strong evidence for this ranking of repeatabilities within most 445 years, and hence in the grand mean across years, in both sexes (OSM S5). Individuals' 446 447 migratory tactics were therefore notably highly repeatable, and the most flexible tactic 448 was the least repeatable (hence, most supraflexible), as expected given the threshold 449 trait model (fig. 2A).

Further, tactic switching was clearly structured along the R-M continuum. 450 Specifically, full-R and full-M individuals switched to mixed-RM more than to full-M 451 452 and full-R respectively (fig. 5; OSM S5). Grand mean switching probabilities were 0.10 [0.07,0.14] and 0.11 [0.08,0.14] from full-R to mixed-RM, and 0.24 [0.19,0.28] and 453 454 0.22 [0.17,0.26] from full-M to mixed-RM, in females and males respectively. In 455 contrast, these probabilities were 0.04 [0.02,0.06] and 0.02 [0.01,0.04] from full-R to full-M, and 0.04 [0.02,0.07] and 0.05 [0.2,0.07] from full-M to full-R. Meanwhile, 456 457 mixed-RM individuals switched in higher and roughly similar proportions to both inflexible tactics, with grand mean probabilities of 0.18 [0.13,0.24] and 0.19 [0.14,0.23] 458 459 for switching to full-R, and 0.28 [0.22,0.34] and 0.15 [0.11,0.19] for switching to full-M, in females and males respectively (fig. 5; OSM S5). These relative probabilities of 460 different forms of supraflexibility again qualitatively match expectations given the 461 462 threshold trait model (fig. 2B).

463	Probabilities of tactic switching also varied among years (fig. 5; OSM S5).
464	Interestingly, full-R females and males were especially likely to switch to another tactic
465	in the two consecutive ECE years for which our model could detect coincident
466	phenotypic changes (2012-13 and 2013-14, fig. 5). However, there was no clear
467	evidence of a difference between these years, or with several non-ECE years (2011-12
468	and 2017-18 in females, 2014-15 and 2018-19 in males; fig. 5; OSM S5). Large among-
469	year variation in switching probabilities also occurred for other tactics. Notably, mixed-
470	RM showed strong evidence for changes in the magnitude and direction of switching
471	in both sexes (OSM S5).
472	Finally, in several years, there was some or strong evidence for differences in
473	switching probabilities between males and females (fig. 5, OSM S5). For example,
474	females were often more likely to switch from mixed-RM to full-M than males, while
475	full-M males showed noticeably consistent switching probabilities compared to females
476	(fig. 5, OSM S5). Supraflexibility was therefore somewhat sexually dimorphic.
477	
478	Survival selection on tactics and flexibility
479	Our analyses revealed notable variation in the shape and strength of survival selection
480	on the defined migratory tactics, and hence on within-year flexibility, between the
481	pooled non-ECE years and each ECE year, and between females and males (fig. 6).
482	In non-ECE years, survival probabilities were very high in both sexes (fig. 6).
483	Nonetheless, in males, there was strong evidence for disruptive selection along the R-
484	M continuum, such that full-R and full-M had higher survival probabilities than mixed-
485	RM ( $\Delta$ : 0.06 [0.00,0.13] and 0.08 [0.01,0.15], respectively; fig. 6). In contrast, there

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486 was no evidence that survival differed among tactics in females ( $\Delta$ : 0.01 [-0.06,0.08] 487 and 0.03 [-0.05,0.12]; fig. 6).

488 In 2012-13, the first ECE year, there was strong evidence for directional selection 489 towards full-M along the R-M continuum in males, and considerable evidence for a similar pattern in females (fig. 6). In both sexes, survival probability was higher for 490 491 full-M than mixed-RM ( $\Delta$ : 0.14 [-0.01,0.29] in females, 0.16 [0.00,0.32] in males; fig. 492 6), and even higher than for full-R ( $\Delta$ : 0.20 [0.08,0.32] in females, 0.33 [0.21,0.46] in males). However, survival probability was lower for full-R than mixed-RM in males 493 494  $(\Delta: -0.17 [-0.31, -0.03])$ , but this was not evident in females ( $\Delta: -0.06 [-0.22, 0.10]$ ; fig. 495 6).

In 2013-14, there was strong evidence for disruptive selection in females (fig. 6), such that both full-R and full-M had higher survival probability than mixed-RM ( $\Delta$ : 0.31 [0.09,0.52] and 0.18 [-0.02,0.38], respectively; fig. 6). There was also some evidence of additional skew in the shape of selection, with higher survival probability in full-R than full-M ( $\Delta$ : 0.13 [-0.04,0.29]; fig. 6). In contrast, in males there was more evidence for stabilising selection than for the other possible shapes of selection, but no strong evidence that survival probability differed between tactics (fig. 6).

In 2017-2018, there was some evidence for directional selection towards full-M in females, and for stabilising selection in males (fig. 6). More precisely, in both sexes, there was some or strong evidence that full-R had lower survival probability than mixed-RM ( $\Delta$ : -0.14 [-0.30,0.04] in females, -0.21 [-0.33,-0.06] in males; fig. 6) and full-M ( $\Delta$ : -0.22 [-0.34,-0.10] in females, -0.13 [-0.25,-0.02] in males). However, there was no clear evidence for differences in survival probability between full-M and mixed-RM ( $\Delta$ : 0.08 [-0.10,0.29] in females, -0.07 [-0.22,0.11] in males; fig. 6).

# Net effects on phenotypic dynamics

The three sequential processes of selection, supraflexibility and entry of new 512 513 individuals each caused detectable immediate changes in tactic frequencies, and hence in the degree of within-winter migratory flexibility, across multiple years in each sex 514 515 (fig. 7). On average across all years, supraflexibility caused the largest changes in tactic frequencies, while survival selection and entry of individuals each caused changes of 516 517 smaller but similar sizes (fig. 7; OSM S5). Yet, these changes varied considerably in 518 size and direction among years (fig. 7; OSM S5). Within-generation effects of survival selection on phenotype frequencies were consequently either accentuated or counter-519 520 acted (or even cancelled) by subsequent supraflexibility in different instances. For 521 example, the reduced frequency of full-R associated with the 2012-13 ECE resulted from both supraflexibility (between year 2011-12 and 2012-13) and survival selection 522 523 (through 2012-13), but was quickly reversed across subsequent years (fig. 7; OSM S5).

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# Discussion

Within-individual and among-individual variation in phenotype, and associated variation in fitness components, underpins joint phenotypic, population dynamic and micro-evolutionary responses to environmental variation and change (Nussey et al. 2008; Chevin et al. 2010; Coulson et al. 2010). Yet, hierarchical patterns of individual phenotypic flexibility within and across biologically-relevant time periods, and associated episodes of directional, stabilising or disruptive selection affecting flexibility, are rarely quantified. Our analyses of ten years of large-scale band-

533 resighting data revealed substantial among-individual variation and structure in flexibility and supraflexibility of seasonal migration versus residence, a key trait that 534 directly shapes spatio-seasonal population dynamics, in European shags. We also 535 536 demonstrate forms of disruptive and directional survival selection on migratory tactics along the residence-migration continuum, that varied among years and occasionally 537 between the sexes. We thus illustrate that hierarchical variation in flexibility across 538 539 timescales in dichotomous phenotypes can be structured in accordance with principles of threshold trait expression and can be subject to complex and varying selection, 540 541 thereby shaping population-wide phenotypic dynamics.

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#### Hierarchical variation in flexibility

Our focal adult shags showed substantial among-individual variation yet notably high within-individual repeatability in annual tactics, including high repeatability of both seasonally inflexible phenotypes (full-R and full-M) and seasonally flexible phenotypes (mixed-RM, fig. 4, 5). This implies that both inflexibility and flexibility in the focal threshold trait could be underpinned by substantial genetic and/or permanent environmental effects.

To our knowledge, four previous studies have quantified within-individual variation in plasticity across different time periods and/or environmental axes alongside among-individual variation (Araya-Ajoy and Dingemanse 2017; Mitchell and Biro 2017; Cornwell et al. 2019; Strickland et al. 2021). All four studies considered continuously distributed behavioural traits and quantified phenotypic variation as reaction norms slopes given known environmental variables, finding considerable repeatability through time. In general, such repeatability could readily occur if plastic 557 responses are repeated because forms of environmental variation that drive reaction norms are repeated through time, for example due to environmental cyclicity such as 558 seasonality (Schlichting and Pigliucci 1998; Piersma and Drent 2003; Lande et al. 559 560 2019). Any underlying permanent environmental effects might reflect habituation or social learning for behaviours including migration (Chapman et al. 2011; Jesmer et al. 561 2018). Repeatability in phenotypic plasticity could also occur in non-repeated 562 563 environmental contexts if individual reaction norm slopes are consistent across different environmental axes (e.g. Cornwell et al., 2019), as expected if common 564 565 physiological or cognitive machineries underpin plasticity in multiple contexts 566 (Coppens et al. 2010; Dingemanse and Wolf 2013; Stamps 2016). Accordingly, withinindividual phenotypic variation has been interpreted to result from responses to internal 567 or external variables that vary through time at multiple scales within a population 568 569 (Araya-Ajoy and Dingemanse 2017; Mitchell and Biro 2017; Cornwell et al. 2019; Strickland et al. 2021), representing multidimensional plasticity (Westneat et al. 2015, 570 571 2019).

In contrast, the threshold trait model, where an underlying environmentally-572 sensitive liability leads to discrete phenotypic outcomes (fig. 1), can directly predict 573 574 structured patterns of hierarchical variation and repeatability in flexibility across timescales without necessarily requiring multidimensional plasticity. Specifically, 575 because the degree of individual phenotypic flexibility is expected to decrease with 576 577 increasing distance of initial liability from the threshold (i.e. with liability elevation, Reid and Acker 2022), among-individual variation in flexibility with directional 578 579 supraflexibility and positive covariation between flexibility and supraflexibility (fig. 1, 580 2) can readily arise, as observed for migration versus residence in shags (fig. 4, 5).

581 Some studies on migration versus residence in other taxa also provide evidence of among-individual variation in within-year flexibility, but did not explicitly consider 582 within-individual repeatability of such flexibility (e.g. roe deer, Capreolus capreolus: 583 584 Cagnacci et al. 2011; blackbird, Turdus merula: Fudickar et al. 2013; red deer, Cervus elaphus: Peters et al. 2019). Other studies demonstrate high repeatability of migration 585 586 versus residence with some between-year flexibility, but did not explicitly consider 587 within-year flexibility (e.g. American dipper, Cinclus mexicanus: Gillis et al. 2008; redspotted newt: Notophtalmus viridescens: Grayson et al. 2011; skylark, Alauda arvensis: 588 589 Hegemann et al. 2015, noctule bat Nyctalus noctule: Lehnert et al. 2018; various 590 ungulates: Berg et al. 2019). Following our current analyses, future studies on these and 591 other systems could explicitly quantify structured patterns of hierarchical variation in 592 phenotypic flexibility, and hence assess general accordance with the threshold trait 593 model. Environmental variables known to be associated with cross-sectional changes 594 in migratory proportions (e.g. individual condition, conspecific density, climate 595 variables, or resource availability; Brodersen et al. 2008; Skov et al. 2010; Grayson and 596 Wilbur 2009; Boyle et al. 2010; Eggeman et al. 2016) would then be relevant candidates 597 to explicitly consider sex-specific liability-scale reaction norms underlying phenotypic flexibility and supraflexibility and hence formally link concepts of observed-scale 598 599 flexibility and latent liability-scale plasticity.

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## Selection on flexibility

602 Our analyses also demonstrate occurrences of notably strong survival selection on 603 annual migratory tactics and hence on migratory flexibility, with substantial variation 604 in the shape and strength of selection between non-extreme years and each ECE year, 605 and sometimes between females and males (fig. 6). Selection against full-R during the ECE years 2012-13 and 2017-18 reflects greater exposure of resident versus migrant 606 607 locations to extreme late-winter storms that caused mortality in these two years (broadly 608 consistent with previous cross-sectional analyses that did not consider annual migratory 609 tactics; Acker et al. 2021a). The lower survival probability of mixed-RM compared to 610 full-M in 2012-13, generating overall directional selection towards full-M, could then 611 effectively represent a cost of flexibility (here, of moving between locations during 612 winter) that prevented the full potential benefit of migration. The disruptive selection 613 along the R-M continuum that was evident for the five non-extreme years in males, and 614 in one ECE year in females (2013-14; fig. 6), is also consistent with costs of mixed-615 RM. In general, such costs might result from fine spatio-temporal variation in survival 616 if mixed-RM individuals experience temporarily poor conditions both before and after 617 departure. They could also represent direct costs resulting from reduced local site familiarity or increased susceptibility to local environmental deteriorations compared 618 619 with individuals that winter in a single location (Skov et al. 2010), and/or from maintaining physiological machineries required to respond to environmental variation 620 621 (e.g. DeWitt et al. 1997; Auld et al. 2010).

622 Still relatively few studies have quantified variation in fitness components in relation to phenotypic plasticity in any trait, and very few considered labile traits. 623 Conclusive evidence of selection on phenotypic plasticity, or on resulting phenotypic 624 625 flexibility, consequently remains scant (Van Buskirk and Steiner 2009; Chevin et al. 2010; Arnold et al. 2019). A general meta-analysis concluded that fitness costs and 626 627 benefits of developmental plasticity were usually small, but that costs predominantly arise under stressful environmental conditions (Van Buskirk and Steiner 2009). In 628 629 contrast, our results indicate that costs of flexibility (i.e. lower survival of mixed-RM,

resulting in disruptive selection) can be manifested under both benign (non-ECE) and
extreme (ECE) conditions (fig. 6). Yet, our results also show that any such costs are
highly contingent, and may often be hidden by large variation in sex- and yeardependent costs and benefits of the alternative inflexible phenotypic tactics (i.e. full-M
and full-R).

635 Conversely, there was some evidence of stabilising selection, defined by higher 636 survival probability of the flexible mixed-RM tactic than both inflexible tactics, only in 2013-14 and 2017-18 in males. There was consequently no strong evidence that the 637 relatively high observed frequency of migratory flexibility could be directly maintained 638 639 through survival selection. However, the threshold trait model (fig. 1) implies that flexibility could be induced by instances of phenotypically disassortative mating. This 640 can arise because any reproduction involving two parents with opposite inflexible 641 642 phenotypes, given additive genetics effects on their liabilities, can readily generate offspring with liabilities close to the threshold that are consequently highly 643 644 phenotypically flexible (Reid and Acker 2022). Variation in phenotypic flexibility in a threshold trait might then be maintained as an indirect consequence of disruptive 645 646 selection that acts to maintain both inflexible phenotypes.

647 Indeed, under the threshold trait model, any form of selection along the continuum of observed phenotypic tactics implies some similar form of selection along 648 649 the underlying liability (fig. 1,2; Reid and Acker 2022). While substantial additive genetic variation in liabilities can be maintained even given consistent directional 650 651 selection, phenotypic variation is unlikely to be maintained without some additional form of non-linear or varying selection (Roff 1996, 1998). These requirements could 652 be effectively fulfilled by the disruptive selection observed in non-ECE years, and the 653 654 variation in the shape of selection observed across ECE years, in turn implying that

655 increasing frequencies of diverse ECEs could help maintain genetic variation in flexibility. Yet, the frequency of full-R typically exceeded other phenotypic tactics, 656 despite having no clear survival advantage across years. This may partly be explained 657 by reproductive selection, since our previous cross-sectional analyses showed that 658 shags that were resident on Isle of May during late winter commonly had higher 659 breeding success than migrants that spent the late winter elsewhere (Grist et al. 2017; 660 661 Acker et al. 2021b). Such relationships could represent carry-over effects of winter location on subsequent reproductive success, generating indirect selection on migratory 662 663 tactic.

664 Meanwhile, few or no previous studies on any trait have considered whether selection on phenotypic plasticity or flexibility is sex-specific (Van Buskirk and Steiner 665 666 2009; Arnold et al. 2019; but see Fox et al. 2019). We observed a striking difference in 667 the form of selection between females and males in the ECE year 2013-14, with apparently strong selection against migratory flexibility in females but not males. There 668 669 was also some weak evidence of selection for flexibility in males but not females in 2017-18. These effects could reflect interactions between environmental variation and 670 671 sex-specific environmental sensitivity in shags. For example, females' foraging 672 efficiency is more sensitive to strong winds, and females can experience higher associated mortality (Lewis et al. 2015; Acker et al. 2021a). Given the threshold trait model, such 673 sex-specific selection on flexibility could translate into sex-specific selection on mean 674 675 liability, due to the expected intrinsic relationship between the two. Thus, while sexspecific selection on any continuously distributed trait could cause or accentuate sexual 676 677 dimorphism (Kruuk et al. 2008; Cox and Calsbeek 2009), such effects could be exacerbated in a threshold trait if sex-specific selection affects flexibility. However, 678 679 since there was no strong sexual dimorphism in the shape of selection in most years,

any such evolutionary consequences are likely to be minor in our system, unless ECEsbecome more frequent.

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# Overall implications for phenotypic dynamics

684 One general yet rarely achieved ambition in evolutionary ecology is to partition effects of multiple simultaneously or sequentially acting processes on overall phenotypic 685 686 variation and change (e.g. Charmantier et al. 2008; Chevin et al. 2010; Coulson et al. 687 2010; de Villemereuil et al. 2020). Our analyses allowed evaluation of the relative 688 contributions of supraflexibility, survival selection on flexibility and entry of new adult individuals in shaping population-wide phenotypic change among adults across the 689 690 study period, including change in flexibility. Despite the strong observed episodes of 691 selection, our results demonstrate some predominance of supraflexibility over selection 692 on flexibility in generating immediate changes in tactic frequencies, and hence in shaping population-wide occurrences of flexible seasonal migration and resulting 693 694 spatio-seasonal dynamics (fig. 7). Yet, selection could still shape longer-term micro-695 evolutionary and phenotypic dynamics by altering within-generation distributions of 696 genetic values for liability-scale reaction norm elevations and/or slopes and 697 corresponding frequencies of flexible and inflexible tactics, and of associated 698 supraflexibility, in subsequent generations. Such selection could partly account for the 699 observed effects of entry of new adults, which are primarily new recruits to the local 700 breeding population, and hence represent contributions from successive (overlapping) 701 generations.

The full longer-term dynamics of phenotypes, and associated degrees of
 flexibility and supraflexibility, will consequently depend on additive genetic variances

704 in liability for migration versus residence and patterns of (dis)assortative mating, 705 alongside developmental plasticity and selection occurring through reproduction and sub-adult survival to recruitment (Charlesworth 1994; Kruuk 2004; Coulson et al. 2010; 706 707 Reid and Acker 2022). These dynamics could be further modulated by complex forms 708 of direct and indirect selection on supraflexibility, including indirect selection resulting 709 from the expected positive covariation with flexibility (fig. 2C). Our accumulating 710 dataset will allow future estimation of all these components, but this will require further 711 major methodological developments. New multistate capture-recapture model 712 structures will be required to appropriately represent memory of the previous phenotypic tactic, reproduction, early-life histories up to recruitment, and phenotype-713 714 dependent mating dynamics. Combining capture-recapture and generalized linear 715 mixed model approaches will ultimately allow estimation of effects on the latent 716 liability scale, including plasticity reaction norms and additive genetic variances. 717 Together, these advances will allow full dissection of all processes shaping short-term and longer-term changes in migratory phenotypes and migratory flexibility, and 718 719 resulting spatio-seasonal population dynamics.

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731	
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733	PA and JMR formulated the ideas and conceptual developments. JMR and FD
734	conceived the migration field study. FD, SW, MPH, MAN, SJB, CG, RS and JMR
735	organised and undertook long-term field data collection. PA designed the modelling
736	methodology, coded the models, analysed the data and drafted the manuscript, assisted
737	by JMR. FD, SJB, and APP contributed manuscript edits. All authors gave final
738	approval for publication.
739	
740	Data and code accessibility
741	Data and code are available from the Dryad Digital Repository:
742	https://doi.org/10.5061/dryad.3j9kd51mn.
743	
744	Literature cited
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Figure 1. Basic principles of the threshold trait model. (A) Each individual expresses 943 944 alternative phenotype X or Y (light grey versus dark grey shading) when its underlying 945 liability is below or above the threshold (dashed line). Liability is conceptualised as a continuous latent variable which can vary among and within individuals due to genetic 946 947 and environmental effects. (B) Hypothetical changes in individual liability through 948 time, given among-individual variation in mean liability (i.e. elevation) and identical 949 time-varying environmental effects. The red and orange individuals have liabilities that 950 remain below and above the threshold (dashed line) respectively, and hence inflexibly 951 express phenotypes X and Y respectively throughout the observed time period (i.e. full-952 X and full-Y tactics). The rose individual has a liability that crosses the threshold at 953 some time, and hence expresses both X and Y (i.e. mixed-XY tactic, representing phenotypic flexibility). (C) Relationship between the degree of individual flexibility 954 955 (i.e. within-individual phenotypic variance) and phenotypic tactic along the X-Y 956 continuum (i.e. expressed proportion of X vs. Y, from full-X to full-Y). Flexibility is zero at both extremes and maximal for a mixed-XY tactic with 50% X and 50% Y (as 957 958 illustrated).



960 961 962 963

selection. (A) Among-individual and within-individual variation in phenotypic tactic, illustrating variation in flexibility and supraflexibility. Phenotypes X and/or Y are expressed during focal time periods (denoted 1 and 2) that occur sequentially. Within each period, each individual expresses either an inflexible phenotypic tactic (full-X in 964 965 red, or full-Y in orange) or a flexible tactic (mixed-XY in rose). There are nine possible 966 paths across two consecutive periods and hence forms of supraflexibility (arrows; OSM S1): individuals can either retain the same tactic (3 possible paths, solid arrows), or 967 968 switch tactic (6 possible paths, dashed arrows). Different arrow widths represent 969 different switching probabilities as qualitatively expected under the threshold trait 970 model, showing that supraflexibility could be structured and covary with flexibility. (B 971 and C) Possible shapes of selection on tactics (and hence flexibility) and supraflexibility

972	arising due to associated variation in fitness ( $\omega$ ), resulting in selection along the X–Y
973	continuum. Such selection could be (B-i) stabilizing ( $\omega_{\text{full-X}} < \omega_{\text{mixed-XY}} > \omega_{\text{full-Y}}$ ) or (B-ii)
974	disruptive ( $\omega_{\text{full-X}} > \omega_{\text{mixed-XY}} < \omega_{\text{full-Y}}$ ) due to relative benefits or costs of being flexible,
975	or directional towards (B-iii) full-X (\alpha_{full-X}>\alpha_{mixed-XY}>\alpha_{full-Y}) or (B-iv) full-Y (\alpha_{full-X})
976	$x < \omega_{mixed-XY} < \omega_{full-Y}$ ) due to relative benefits of expressing X or Y. Further, there could
977	be selection (C-i) for or (C-ii) against supraflexibility depending on fitness variation
978	occurring in two consecutive periods. Such selection on supraflexibility could
979	intrinsically result from sequences of selection on flexibility in consecutive years,
980	and/or from direct costs or benefits of switching tactic between years. It could therefore
981	be partly conditional on, but also partly independent of, the tactics from and to which
982	switching occurs.



Figure 3. Summary of the state transition process of the full-annual-cycle capture-984 recapture model representing a dynamic finite mixture of annual migratory tactics, and 985 986 hence individual variation in flexibility and supraflexibility of migration versus residence. The diagram represents possible fates of individuals from the year they enter 987 988 the dataset (y) to the next year (y+1). An individual undertakes one possible initial tactic (dashed arrows): either full residence ('full-R', top, red), mixed residence-migration 989 990 ('mixed-RM', middle, rose), or full migration ('full-M', bottom, orange). All 991 individuals start the annual cycle during the breeding season at the residency area (state 'B'). Each tactic is defined by constraints that determine possible within-year 992 transitions (white arrows) of surviving individuals between the two phenotypic states: 993 994 residence (R) and migration (M). Then, between years (here year y to y+1), individuals 995 can switch between tactics or repeat the same tactic (coloured arrows). The complete 996 fate diagram showing the full state transition and observation processes is in OSM S3.



**Figure 4**. Relative tactic frequencies estimated at the start of each biological year (i.e. occasion 1;  $F_1$ ) from 2010-11 to 2018-19 in (A) females and (B) males, derived for full residence ('full-R', red circles) and full migration ('full-M', orange triangles), representing tactics that are phenotypically inflexible within the focal winter, and mixed residence-migration ('mixed-RM', rose squares), representing a phenotypically flexible tactic within the focal winter. Point estimates are posterior means, inner and outer line segments indicate 50% and 95% credible intervals.



Figure 5. Probabilities of between-year switching ( $\kappa$ ) from (A) full residence ('full-R'), 1006 1007 (B) mixed residence-migration ('mixed-RM') and (C) full migration ('full-M') to each 1008 of these tactics, in females (top row) and males (bottom row). Probabilities of switching 1009 to full-R, mixed-RM and full-M are in red circles, rose squares and orange triangles 1010 respectively. The probability of switching from any one tactic to the same tactic 1011 represents the probability of repeating this tactic (i.e. not switching). Point estimates 1012 are posterior means, inner and outer line segments indicate 50% and 95% credible 1013 intervals. The x-axis indicates the biological year to which switching occurs (e.g. 2012-1014 13 is for the transition from 2011-12 to 2012-13).



1016 Figure 6. Tactic-specific annual survival probabilities, and associated evidence for 1017 possible shapes of selection, in (A) females and (B) males. Annual survival probability 1018  $(\Phi)$  was estimated separately for each of the three years with an extreme climatic event 1019 ('ECE'), and estimated across all five non-ECE years combined. In the survival 1020 probability graphs, point estimates are posterior means for full residents ('full-R', red 1021 circles), mixed resident-migrants ('mixed-RM', rose squares), and full migrants ('full-1022 M', orange triangles). Inner and outer line segments indicate 50% and 95% credible 1023 intervals. Segment annotations indicate the probability that the difference  $\Delta$ , between 1024 the two points at the left and right end of the segment, is positive ( $Pr(\Delta > 0)$ ; values close 1025 to 0 or 1 indicate substantial evidence for a negative or positive difference respectively; values close to 0.5 indicate no clear evidence for either). In the heatmaps, values are 1026 1027 the probabilities of each possible shape of selection along the residence-migration 1028 continuum (Pr(shape); i: stabilising, ii: disruptive, iii: directional towards full-R, and 1029 iv: directional towards full-M).



Figure 7. Overall synthesis of phenotypic dynamics of the three defined migratory tactics. The walkthrough example (A) illustrates that relative tactic frequency (F) is quantified at the start of each year y once new individuals have entered the adult population ( $F_{1,y}$ , filled coloured circles), then immediately after annual survival selection ( $F_{2,y}$ , open coloured squares), and immediately after subsequent tactic switching between y and y+1 ( $F_{3,y}$ , filled coloured triangles). The corresponding net

1037 effects on variation in tactic frequency (E) from y to y+1, due to survival selection 1038  $(E_{1,y}=F_{2,y}-F_{1,y}, \text{ open black squares})$ , supraflexibility  $(E_{2,y}=F_{3,y}-F_{2,y}, \text{ black triangles})$ , and entry of new individuals ( $E_{3,y}=F_{1,y+1}-F_{3,y}$ , circles) are negative when frequency 1039 1040 decreases and positive when it decreases. Estimates are shown for (B) females and (C) 1041 males, for full residence ('full-R', red, top row); mixed residence-migration ('mixed-1042 RM', rose, middle row); full migration ('full-M', orange, bottom row). Coloured points are posterior means of F, with inner and outer line segments indicating 50% and 95% 1043 1044 credible intervals. Black points are posterior means of E, with lines indicating 90% 1045 credible interval. A 90% credible interval located below (or above) zero indicates that 1046 the probability of *E* being negative (or positive) is >95%.