# 1 Total daily energy expenditure has declined over the last 3 decades due to

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2 declining basal expenditure not reduced activity expenditure.

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149	Abstract	
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151	Obesity is caused by prolon	ged positive energy balance <sup>1,2</sup> . Whether reduced energy
152	expenditure stemming from	n reduced activity levels contributes, is debated <sup>3,4</sup> . Here we show
153	that in both sexes, total ene	ergy expenditure (TEE) adjusted for body composition and age
154	declined since 1981, while a	adjusted AEE increased over time. We use the International Atomic
155	Energy Agency Doubly Labe	lled Water (IAEA DLW) database on energy expenditure of adults in
156	the USA and Europe (n = 47	99) to explore patterns in total (TEE: n=4799), basal (BEE: n = 1432)
157	and physical activity energy	expenditure (AEE: n = 1432) over time. In males, adjusted BEE
158	decreased significantly, but	in females this didn't reach significance. A larger dataset of basal
159	metabolic rate (BMR equiva	alent to BEE) measurements of 9912 adults across 163 studies
160	spanning 100 years replicat	es the decline in BEE in both sexes. We conclude that increasing
161	obesity in the USA/Europe l	nas probably not been fuelled by reduced physical activity leading to
162	lowered TEE. We identify he	ere a decline in adjusted BEE, as a previously unrecognized factor.

Main text: Obesity is a global health threat<sup>5</sup>. Although excess body fat is a result of prolonged 163 164 positive energy balance<sup>1,2</sup>, the exact causes of this imbalance remain elusive. Two major potential factors have been suggested. First, food consumption (net energy consumption 165 accounting for losses in feces) may have increased<sup>2</sup>. Alternatively, declines in energy 166 expenditure, due to reduced work-time physical activity<sup>4</sup>, combined with increases in sedentary 167 behavior, partly linked to elevated 'screen time' (TV, computer and phone use)<sup>6,7</sup> may be a key 168 driver. These may be linked in a vicious cycle<sup>8</sup>, where low activity leads to weight gain, which 169 170 inhibits activity, leading to further weight gain.

171 Although there is direct evidence, that physical activity levels have declined and sedentary time has increased<sup>4,6,7,8</sup>, these changes do not necessarily translate into alterations in 172 173 total energy expenditure (TEE). As individuals get larger the energy cost of movement also increases<sup>9</sup>. Thus, the same amount of energy may be utilized even though the actual time spent 174 175 active has declined. Moreover, increases in one type of activity/behavior may be replaced by decreases in another behavior of equal cost. Consequently, apparently large behavior changes 176 may result in only minor alterations in expenditure. Finally, it has been suggested that we may 177 compensate for changes in physical activity by adjusting expenditure on other physiological 178 tasks<sup>10,11</sup>. Although low TEE is repeatable, and having low TEE is not a risk factor for future 179 weight gain over short timescales<sup>12</sup>, this does not negate a possible impact over longer periods. 180 181 In the present paper we address the idea that reduced physical activity leading to reduced activity energy expenditure (AEE) may have fueled the epidemic. 182

183 The doubly-labelled water (DLW) method is a validated isotope based methodology for 184 the measurement of free-living energy demands<sup>13</sup>. A previous analysis using this method

suggested there had been no change in TEE between 1986 and 2005, calling into question the
reduced physical activity hypothesis<sup>14</sup>. Nevertheless, these observations were based on a
limited sample (n = 314) from a single European city over a restricted timespan of about 20
years. Here we expanded this analysis using data for 4799 adults living across Europe and the
USA drawn from the IAEA DLW database<sup>15</sup> for which we also had BEE measures in 1429
individuals. All estimates of TEE were recalculated using a common equation<sup>16</sup> that has been
shown to perform best in validation studies<sup>16</sup>.

192 We split the data by sex, because this may affect the etiology of energy balance<sup>17,18</sup>. This resulted in 1672 measurements of males and 3127 measurements of females. In addition, for 193 194 632 of the males and 800 of the females we also had measurements of basal energy 195 expenditure (BEE) from which we derived activity energy expenditure (AEE) and physical 196 activity level (PAL) – for calculations see methods. The data span a period of over 30 years with 197 the first measurements in late 1981 and the latest measurements made in late 2017, with most 198 data obtained between 1990 and 2017. The distribution of BMI in the sample for both males and females is shown in Extended data Fig 1. Overall females had higher BMI than males. In the 199 200 pooled sample the distribution was BMI < 18.5 = 2.3%, BMI 18.5 to 25 = 40.3%, BMI 25 to 30 =35.1% and BMI >30 = 22.2%. Combined overweight and obesity was 57.3%. In both males and 201 202 females body weight increased over time (Figure S1) reflecting the secular trend in body weight 203 over the same interval.

We first explored the changes in the unadjusted levels of TEE, BEE and AEE over time (Extended data Fig 2: Table 1). In males there was no significant relationship between TEE and the date of measurement (date coded as months since Jan 1982) (r<sup>2</sup> = 0.0015, p = 0.14 (ns): Fig

S2a) the least squares regression fit gave a gradient of  $\pm 1.5 \text{ kJ/month}$  (95%CI =  $\pm 2.06$ 207 208 kJ/month). This gradient leads to an estimated change in average TEE over 30 years of + 0.55 MJ/day (95%CI = ±0.727 MJ/day). Contrasting the lack of significant change in TEE, there was a 209 significant decline in BEE over time (Fig S2b) ( $r^2 = 0.029$ , p = 0.000018). The gradient of decline 210 211 (3.3 kJ/month, 95%CI = ±1.4 kJ/month) was equivalent to an average fall in BEE by 1.19 MJ (9.7%) over 30 years (95%CI = ±0.54 MJ/day). As might be anticipated since TEE\*0.9 = BEE + 212 AEE, the absence of a change in TEE and declining BEE was reflected by an increase in AEE over 213 214 time, but this did not reach significance (Fig S2c) ( $r^2 = 0.003$ , p = 0.16). The gradient of the 215 change in AEE (1.4 kJ/month, 95%Cl =  $\pm 1.8$  kJ/month) was equivalent over 30 years to an increase by 0.50 MJ/day (95%CI = ±0.69 MJ/day). In females, unadjusted levels of TEE, BEE and 216 217 AEE did not change significantly over time (Extended data Fig 3, Table 1). All the energy expenditure variables (TEE, BEE and AEE) in both sexes were dependent 218 219 on body mass (BM) and BMI (illustrated for BMI in Extended data Fig 4). Because of these relationships it is necessary to adjust the raw expenditure data over time (Extended data Figs 2 220 221 and 3) to account for any changes in body composition over time that might generate a biased

estimate of the change in expenditure variables. We adjusted the levels of log transformed TEE,

223 BEE and AEE for body size and composition using residuals from general linear models with loge

fat-free mass, loge fat mass and age as predictors. In this analysis the data were logged because

the relationships between energy expenditure components and body composition follow

power law relationships. In males, adjusted TEE significantly declined over the measurement

period (Fig 2a:  $r^2$  = 0.0103, p < .0001). The gradient of the fitted regression was -2.58 kJ/month

228 (95%CI =  $\pm 1.20$  kJ/month) leading to an estimated average change over 30 years of -0.93

MJ/day in adjusted TEE (95%CI = ±0.465 MJ/day), a decline on average of 7.7%. The adjusted 229 230 BEE showed a highly significant decline over time (Fig 2b:  $r^2 = 0.064$ ,  $p < 10^{-9}$ ) with the gradient of -2.67 kJ/month (95%CI = ±0.82 kJ/month) being equivalent to an average fall in BEE of 0.96 231 232 MJ/day (14.7%) over 30 years (95%CI = ±0.15 MJ/day). In contrast, the adjusted AEE increased over time (Fig 2c:  $r^2$  = 0.0221, p < .0003). The gradient of +2.8 kJ/month (95%Cl =  $\pm 1.4$ 233 kJ/month) was equivalent to a rise of 1.01 MJ/day over 30 years (95%CI = ±0.53 MJ/day). 234 In females as well, there was a significant decline in the adjusted TEE over time (Fig 3a: 235 236  $r^2$  = 0.006, p < .00002). The gradient of the effect 1.42 kJ/month was equivalent to a reduction 237 in TEE over 30 years of 0.51 MJ (95%Cl =  $\pm 0.22$  MJ/day) or 5.6%. This decline was paralleled by a reduction in adjusted BEE of 2.0% but this did not reach significance (Fig 3b:  $r^2 = 0.0015$ , p 238 239 =0.071). The gradient of the fall in adjusted BEE was 0.3 kJ/month, equivalent to a reduction in adjusted BEE over 30 years of 0.11 MJ/day (95%CI =  $\pm 0.21 \text{ MJ/day}$ ). In contrast, and again 240 similarly to the males, adjusted AEE significantly increased over time (Fig 3c:  $r^2 = 0.0063$ , p =241 242 0.026). The gradient of increase in AEE of 1.16 kJ/month was equivalent to an increase in AEE of  $0.42 \text{ MJ/day over 30 years } (95\% \text{ CI} = \pm 0.37 \text{ MJ/day}).$ 243 Because there was a small sample of measures in the early 1980s in males these may 244 have exerted undue leverage in the regression models. We therefore repeated the analysis 245 excluding these data. Their removal had no impact on the detected relationships 246 247 (Supplementary Table S1). Since individual studies may also exert undue leverage we performed additional sensitivity analyses on the BEE effect (post 1987) where the data for each 248 249 study was systematically removed and the regression recalculated. In males removal of no

250 individual study resulted in the loss of significance (Supplementary Table S2). In females

251 however, the absence of significance was due to inclusion of data from a single study

(Supplementary Table S3). We have no reason to exclude these data, but their undue influence
may explain the anomalous lack of decline in female BEE when TEE is declining and AEE is rising
(Table 1 and fig 2).

255 Hence, in both males and females there was a decline in the adjusted TEE by 7.7 and 256 5.6% respectively and in males in the adjusted BEE over time by 14.7% over 30 years (females 257 declined by 2% which was not significant). In both sexes the confidence limits for the decline in 258 adjusted TEE overlapped with the confidence limits for the decline in adjusted BEE, suggesting 259 the decline in adjusted BEE could be sufficient to explain the reduction in adjusted TEE. In both 260 sexes there was in contrast a significant increase over time in adjusted AEE. The comparable 261 declines in adjusted TEE and BEE resulted in a significant increase in PAL (=TEE/BEE) in males (Males Extended data Fig 5a:  $r^2 = 0.0215$ , p < .0003) but in females the change in PAL over time 262 was not significant (females Extended data Fig 5b:  $r^2 = 0.0037$ , p = 0.085). 263

264 To replicate and check our observation of decreasing BEE over time we systematically reviewed data from the literature on mean BMR over the last 100 years, restricted to studies in 265 the USA and Europe, to match the restricted regions included in the time course from the IAEA 266 267 database (Figs 1,2 and Table 1). For the distinction between BEE and BMR see the methods. The 268 main effect on Log<sub>e</sub> BMR was Log<sub>e</sub> BM (Fig 3a), with additional effects of sex and age (total  $r^2$  = 269 0.88). Including the date of measurement, sex, age and loge body mass as predictors in a weighted regression analysis there was a significant negative effect of date of measurement (R<sup>2</sup> 270 271 = 0.024, p = 0.022) on the adjusted  $\log_e$  BMR (Fig 3b). On average, BMR adjusted for BM, age

and sex has declined by about 0.34 MJ/d over the last 100 years. This decline is consistent with,
but at a lower rate, than the data from the IAEA database reported above (Table 1).

Basal metabolism may be influenced by many factors one of which is diet. Human 274 275 dietary changes during the epidemic have included many things such as changes in the amounts 276 of fiber and fat, and the types of fat consumed. Because evaluating the impacts of long-term 277 diets on human metabolism is difficult, we explored the potential impact of dietary fatty acids on metabolic rate using the mouse as a model. Working with mice has the advantage that diets 278 279 can be rigorously controlled and maintained constant over protracted periods. We exposed 280 adult male C57BL/6 mice to 12 diets (for details see supplementary Table S4) that varied in their fatty acid composition for 4 weeks (equivalent to 3.5 years in a human). Mouse BMR (kJ/d) was 281 strongly related to body weight (regression  $r^2 = 0.512$ , p =  $3x10^{-11}$ : Fig 4A). We included the 282 total intake of different fatty acids (SAT: saturated fatty acids, MUFA: mono-unsaturated fatty 283 284 acids and PUFA: poly-unsaturated fatty acids) with body weight into a general linear model. 285 Only intake of saturated fatty acids was significant (SAT: F = 11.05, p = 0.002 (Fig 4B); MUFA: F = 1.38, p = 0.245; PUFA: F = 0.17, p = 0.686) with higher levels of SAT linked to higher energy 286 287 expenditure (Fig 4B).

Overall the data we present do not support the idea that lowered physical activity in general, leading to lowered energy expenditure, has contributed to the obesity epidemic during the last 30 years. Unadjusted AEE was higher in individuals with greater BMI (Extended data Fig 4). This is because, as shown previously, despite on average moving less, individuals with greater BMI have higher costs of movement<sup>9</sup>. Rather than adjusted AEE declining, it has significantly increased overtime in both sexes. Yet TEE (adjusted for age and body composition)

has declined significantly in both males and females over the past 3 decades. Because adjusted 294 295 AEE has increased at the same time that TEE has declined there has been a corresponding 296 reduction in adjusted BEE (which only reached significance in males). The observation that 297 adjusted AEE (and PAL in males) has significantly increased over time is counter intuitive given 298 the patterns established in worktime physical activity and the suggested progressive increase in sedentary behavior<sup>4,6-8</sup>. One possibility is that lowered work time physical activity may have 299 been more than offset by increased engagement in leisure time physical activity. For example, 300 301 sales of home gym equipment in the USA increased from 2.4 to 3.7 Bn US\$ between 1994 and 2017<sup>19</sup>. Time spent in leisure time PA in the USA also increased between 1965 and 1995,<sup>20</sup> 302 suggesting leisure time PA has replaced the decline in worktime PA levels<sup>20</sup>. Leisure time PA has 303 also changed in other westernized populaions<sup>21</sup>. Although time spent on computers has 304 increased, much of the increase in this time has largely come at the expense of time spent 305 watching TV. Since these activities have roughly equivalent energy costs<sup>22</sup> this change would 306 307 not be apparent as a decline in overall adjusted AEE.

The reduction in adjusted BEE is less easily understood but is consistent with the recent 308 observation that body temperatures have also declined over time<sup>23</sup>, over the same interval as 309 the reduction of BMR in the wider data set we analysed (Fig 3b). The magnitude of secular 310 change in BMR is consistent with studies measuring BMR and body temperature in several 311 312 contexts, including calorie restriction, ovulation, and fever which show a 10-25% increase in BMR per 1°C increase in core temperature<sup>24,25</sup>. It was recently suggested that changes in both 313 activity and basal metabolism may have contributed to the decline in body temperature  $(T_b)^{26}$ , 314 but our data suggest this is probably dominated by a BMR effect. The reduction in T<sub>b</sub> has been 315

316 speculated to be a consequence of a reduction in baseline immune function because we have 317 greatly reduced our exposure to many pathogens. However, the links between immune 318 function and metabolism are not straightforward. For example, artificial selection on metabolic rate leads to suppressed innate but not adaptive immune function<sup>27</sup>, and studies of birds point 319 320 to no consistent relation between immune function and metabolism either within or between subjects<sup>28</sup>. Experimental removal of parasites in Cape ground squirrels (*Xerus inauris*) led to 321 elevated rather than reduced resting metabolic rate<sup>29</sup>. Nevertheless, some studies in forager-322 323 horticulturalist societies in South America have noted elevated BMR is linked to increased levels of circulating IgG<sup>30</sup> and cytokines<sup>31</sup>, supporting the view that a long term decline in BEE may be 324 mediated by reduced immune function. Whether this has any relevance to changes in the 325 326 USA/Europe in the past 30 years is unclear. It is also possible that the long-term reduction in BMR represents methodological artefacts. In the early years, measurements of BMR were often 327 328 made using mouthpieces to collect respiratory gases, and recently such devices have been shown to elevate BMR by around 6%<sup>32</sup>. A second possibility is that early measurements paid 329 less attention to controlling ambient temperature to ensure individuals were at thermoneutral 330 temperatures<sup>33</sup>. 331

During the past century there have been enormous changes in the diets of US and European populations (USDA and FAO food supply data)<sup>34</sup>. These changes have included alterations in the intake of carbohydrates, fiber and fats, with % protein intake remaining relatively constant<sup>34</sup>. While intake of carbohydrates peaked in the late 1990s the intake of fat has increased almost linearly since the early part of the 1900s. Moreover, the fat composition has changed dramatically with large increases in soybean oil and seed oils from the 1930s

338 onwards (dominated by the polyunsaturated 18:2 linoleic acid and other PUFAs) and reductions 339 in animal fats (butter and lard) (dominated by saturated fatty acids palmitic (16:0) and stearic acid (18:0) and the mono-unsaturated oleic acid (18:1))<sup>34</sup>. The change has been dramatic, as 340 animal fats comprised >90% of the fatty acid intake in 1910 but are currently less than 15%. 341 342 Because linoleic acid is desaturated to form arachidonic acid (ARA) and ARA is linked to endocannabinoids it has been speculated that expanding linoleic acid in the diet may be linked 343 to various metabolic issues. Effects on basal metabolic rate however are disputed, and if 344 345 anything, PUFAs lead to elevated not reduced metabolism<sup>35,36</sup>, although many studies suggest no effect<sup>37,38</sup>. This variation in outcome may reflect difficulties in controlling human diet over 346 347 protracted periods necessary to generate robust changes in metabolism. In mice, where we can 348 rigorously control the diet for prolonged periods (equivalent to many years of human life), we have shown here no effect of PUFAs on metabolic rate, but a clear impact of saturated fat, with 349 350 greater intake of saturated fat leading to higher metabolic rate (adjusted for body mass). This 351 finding is consistent with earlier reports of relationships between membrane lipids and elevated metabolic rate in mice, particularly a positive effect of palmitic and stearic acids<sup>39,40</sup>. 352 This suggests that alterations in the intake of saturated relative to unsaturated fat over the last 353 354 100 years may have contributed to the decline in BEE reported here, although clearly we should 355 be cautious about extrapolations from males of a single inbred mouse strain and further studies 356 in humans are required. Moreover, other aspects of the diet that impact metabolic rate may also have changed over time, for example intake of fiber which has declined in recent years<sup>41</sup> 357 and has been shown in a randomized controlled trial to affect resting metabolic rate<sup>42</sup>. 358

359 Strengths and limitations

360 A strength of this study is the large sample of individuals over a restricted geographical 361 area (US and Europe) measured using a complex methodology. This has allowed us to detect a 362 small but nevertheless biologically meaningful signal. However, it is important to be aware that the studies were not designed with the current analysis in mind. Hence while we have adjusted 363 for differences in age and body composition there may be other factors that differed over time 364 that we did not adjust for and that could explain the trends we found. Further, the participants 365 366 recruited at different time points may not have been representative of the underlying 367 populations, even though the overall distribution seems representative (Fig S1). The data are cross-sectional which limits the inferences that can be made regarding causality in the 368 369 associations. Finally, while we have speculated on some potential factors that might have 370 contributed to the reduction in BEE (i.e. immune function and diet), these factors were not quantified in most of the participants who had their TEE measured. The mouse work we 371 372 performed showing potential links of diet to metabolism was only conducted in males of one 373 strain and a single age and may not be more broadly applicable. These potential mechanisms therefore remain speculations until more direct data can be collected. 374

# 375 Conclusion

Overall our data show that there has been a significant reduction in adjusted TEE over the last three decades, which can be traced to a decline in BEE rather than any reduction in AEE linked to declining physical activity levels. Indeed, our data show that AEE has significantly increased over time. Reductions in BEE extend much further back in time (TEE data do not extend further back than 1981 as that was the first year the DLW technique was applied to humans), and mouse data indicated that one of many possible explanations is decreases in the

intake of saturated relative to unsaturated fatty acids. If the decline in BEE over time has not
been compensated for by a parallel reduction in net energy intake then the energy surplus
resulting would be deposited as fat. This study therefore identifies a novel potential contributor
to the obesity epidemic, that has not been previously recognized: a decline in adjusted BEE
linked to reduction in overall adjusted TEE. Further understanding the determinants of BEE and
the cause of this decline over time, and if it can be reversed, are important future goals.

388

## 389 Materials and methods

This study involves in part a retrospective analysis of data submitted to the IAEA DLW database (www.dlwdatabase.org). The data stretch back to the late 1980s, however, the clinical trials registry was only launched by the NIH in February 2000, hence, there was no possibility to preregister the work before data collection started. Nevertheless, the analysis performed here was pre-registered on the IAEA DLW database website in 2020 (<u>https://doubly-labelled-water-</u>

395 <u>database.iaea.org/dataAnalysisPlanned</u>).

# 396 DLW database study

<sup>397</sup> Data were extracted from the IAEA Doubly Labeled Water (DLW) Database<sup>15</sup>, version 3.1.2,

compiled in April, 2020, and then later while the manuscript was in review this was expanded

to include additional data extracted from version 3.7.1. In total this latter version of the

400 database comprises 8,313 measurements of TEE using the DLW method. We selected from the

- 401 database measurements of adults aged >18 y, living in either Europe or the USA, that also had a
- 402 record of age. We excluded individuals who were professional athletes, individuals engaged in
- 403 unusual levels of activity (e.g. climbing mountains or participating in a long distance running

race), pregnant and lactating females and individuals with specific disease states. In total this
resulted in 4799 measurements across both sexes. Submissions to the database did not reveal
whether the sex was self-reported or assigned. Although an option was available to designate
individuals as trans-sexual, none of the submitted data were identified as such. Gender was not
available from the submitted data. Estimates of TEE were recalculated using a common
equation<sup>16</sup> which has been shown to perform best in validation studies. The final data set
included 1672 measurements of males and 3127 measurements of females.

411

For 632 of the males and 800 of the females we also had measurements of basal metabolic rate 412 413 (BMR) measured by indirect calorimetry. BMR measurements were derived either from hood 414 calorimetry or from minimal metabolic rate determined overnight during chamber calorimetry (strictly sleeping metabolic rates or SMR). We converted these BMR or SMR to estimates of 415 416 basal energy expenditure (BEE). BMR and SMR are measured for relatively short periods lasting 417 30 minutes to an hour. BEE is a theoretical value for the energy expenditure that would pertain if this BMR/SMR measurement was sustained for 24h. For those individuals with measurements 418 of both BEE and TEE we estimated activity energy expenditure (AEE =  $(0.9^{+}TEE)$ )-BEE), and the 419 420 physical activity level (PAL = TEE/BEE). The value 0.9 in the equation for AEE assumes the 421 thermic effect of food (TEF) is 10% of the total energy expenditure. In practice this varies 422 between individuals and is dependent on the diet. Variation is introduced therefore by 423 imprecision in this value. However, since the TEF is largely dependent on protein in the diet, and protein intakes have remained stable over the last 40 or so years there is unlikely to be any 424 systematic imprecision in the value that could affect the detected trends. It is important to 425

note that TEE and BEE are both measured directly, while AEE is only inferred from the
difference between the two. The accuracy and precision of TEE relative to chamber indirect
calorimetry for the equation utilized here was estimated at 0.4% (accuracy) and 7.7%
(precision)<sup>16</sup>. The accuracy and precision of estimates of basal metabolic rates of metabolism
inferred by indirect calorimetry has been evaluated using alcohol burns and is estimated at
around 1-2%. Error in the estimate of AEE by subtraction is considerably higher than the direct
estimates of TEE and BEE<sup>43</sup>.

433

The DLW method is based on the differential elimination of isotopes of oxygen and hydrogen introduced into the body water<sup>13</sup>. The details of the practical implementation of the method and its theoretical basis have been previously published. We recently derived a new equation for the calculation of CO<sub>2</sub> production using the technique<sup>16</sup> and recalculated the entries in the database using this common equation. These were then converted into energy expenditure using the Weir equation<sup>44</sup> with food quotients derived from the original studies.

440

Additional characteristics of the subjects (body mass (BM), age, and sex) were measured using
standard protocols. We estimated the fat-free mass (FFM) of individuals using the estimated
total body water and an assumed hydration constant for lean tissue of 0.73 (ref 45) and then
calculated fat mass by difference (FM = BM-FFM). The date of the measurement was expressed
in months relative to January 1982 which was the first year that the DLW method was applied
to human subjects.

447

448 In the first set of analyses we used the unadjusted measures of TEE, BEE and AEE as dependent 449 variables in general linear models with time since January 1982 as the predictor. Tests were 450 two-sided and p < .05 was taken as significant. All analyses were performed using Minitab v19. It is well established that TEE, BEE and AEE depend on body composition, as well as subject age. 451 452 Patterns of variation in unadjusted values with time might then reflect biased population 453 sampling with respect to these traits. For example, if more older subjects were sampled later in the time course this might give a spurious indication that TEE was declining since all EE 454 455 parameters decline after ~60y<sup>46</sup>. We adjusted (logarithmically) TEE, BEE and AEE using log<sub>e</sub> FFM, 456 log<sub>e</sub> FM and age as the predictor variables using general linear modelling. Analyses were run 457 separately for each sex therefore no adjustment for sex was necessary. In both sexes, for loge 458 BEE, the predictors age, log<sub>e</sub> FFM and log<sub>e</sub> FM were all significant but for log<sub>e</sub> TEE and log<sub>e</sub> AEE, only age and Loge FFM were significant. In the latter cases we deleted the non-significant 459 460 predictor and re-ran the analyses. Following the above procedure we then calculated the 461 residuals to the fitted models and added them back to the mean logged TEE, BEE and AEE across all measurements. These values were then converted back to measures of 'adjusted 462 TEE', 'adjusted BEE' and 'adjusted AEE' by taking the exponent of the derived values. We then 463 464 checked that the residuals were normally distributed and the adjusted variables were not significantly related to any of the predictor variables to ensure that the adjustment was 465 466 adequate. Tests applied were two-sided and p < .05 was taken as significant. We then sought 467 relationships between the adjusted variables and date of measurement using linear regression. 468 The adjusted variables cover a narrower time span from 1990 to 2017.

469

470 Sensitivity analyses

We performed several checks on the data to make sure the trends were not being driven by
individual studies. First there were some small studies in males prior to 1987 that may have
exerted undue leverage in the analysis. We therefore excluded these data and reran all the
regressions (Table S1). There were no significant changes in any of the parameters. Since the
downward trend in BEE was the most important new finding we directed particular attention to
this trend.

477 To evaluate if the male BEE data would be better fit by a more complex model than the linear model we used, we included higher order terms of the date into a regression analysis. In this 478 analysis the r<sup>2</sup> explained by date, date<sup>2</sup> and date<sup>3</sup> was increased relative to just including date 479 480 alone. However, the variance inflation factors (VIF) for these more complex models were enormous. When date and date<sup>2</sup> were included the VIF for each variable was 28.9, and when all 481 3 were included the VIF values were 438 for date, 2084 for date<sup>2</sup> and 663 for date<sup>3</sup>. The usual 482 483 VIF cut-off for deciding whether to include an extra term into a model is 5. In this case it was clear that higher order terms were not justified relative to a simple linear model. 484 We performed a general linear model analysis with date as a covariate and study as a factor in 485 the model. In males when we used such a model there was indeed a large study effect (F = 486 12.97,  $p < 10^{-15}$ ) but the effect of date remained highly significant (F= 22.87, P <  $10^{-8}$ ) and 487 488 strongly negative (coefficient = -1.85 MJ/d over 30 years), exceeding that in the original 489 analysis. In females there was also a strong study effect (F = 9.54, P < 10-12) but the effect of 490 date remained non-significant (F = 12.9, P = 0.256).

491 Using the post 1987 data we then systematically removed the data for each study and reran the 492 analyses to see if any particular study exerted undue effects on the regression. The analyses are 493 summarized in Table S2. This analysis showed that no individual study was responsible for the negative relationship. In all cases the relationship between BMR and time remained negative 494 495 and highly significant. A single study (number 23 in 1991) involved relatively high BMR values 496 and so omitting it reduced the coefficient and the significance. But the p value for the regression when omitting these data was still highly significant  $P < 10^{-5}$ , and the coefficient still 497 498 strongly negative and biologically important.

We then turned our attention to the female data for BEE against date to see if the absence of a 499 500 relationship there might be because of inclusion of any particular study. We used the same 501 leave one out procedure as used for the males. The results are shown in Table S3. In this case the pattern was very different in that the relationship was always not significant (P > 0.1), 502 503 except when a single study (study 65) was removed from the analysis, and in that case the 504 relationship became significant (P = 0.001) and the negative gradient (extrapolated to over 30 years) increased to -0.39 MJ/day. Omitting a second study (n = 69) has a smaller effect that also 505 resulted in the relationship becoming marginally significant (p = 0.037). If both studies 65 and 506 69 were omitted the p value for the relationship fell to  $P < 10^{-5}$  and the gradient was -0.59 507 MJ/d. Study 65 was a study of overweight individuals<sup>47</sup>. We have no objective reason to reject 508 509 these data but it is interesting that the anomalous absence of a negative relationship of BMR to 510 time in the females is dependent only on inclusion of this one study. It is worth noting that 511 excluding this study from the male data strengthened the relationship for males (Table S2).

512

# 513 Mouse indirect calorimetry measurements

514 All mouse studies followed the guidelines issued by Yale University's Institutional Animal Care and Use Committee (IACUC). Male C57BL/6J mice (Jackson Laboratories, stock # 000664) arrived 515 at the facility at 5 weeks of age and were kept on a 12h/12h light/dark cycle and had free access 516 517 to water and chow diet (Envigo Teklad, 2018S). At 6 weeks of age, mice were switched to one of 518 the different high-fat diets (Research Diets Inc., Table S4). The high-fat diets (HFD) contained 20% protein, 35% carbohydrates and 45% fat by energy with the fat being derived from different 519 520 sources (listed in Table S5). After 4 weeks of HFD feeding, mice were housed in a TSE 521 PhenoMaster system for 4 days. Data from the final 72 hours were used for calculations. Oxygen (O<sub>2</sub>) consumption (mL/h), carbon dioxide (CO<sub>2</sub>) production (mL/h) and food intake (g) were 522 523 recorded every 30 minutes. Energy expenditure (kJ/h) was calculated using the Weir Equation<sup>44</sup>. Respiratory exchange ratio (RER) was calculated as vCO<sub>2</sub>/vO<sub>2</sub>. 524

525

## 526 **Preparation of samples for GCMS**

For mouse diets, approximately 40-50 mg of pulverized diet was weighed and dissolved in 0.5 mL 527 of pure water, acidified with 10 µL of 1 M HCl, and 1 mL of 100% methanol was added. Diet 528 samples were mechanically homogenized to a uniform slurry. Total lipid extraction was 529 performed on all samples as previously described<sup>48</sup>. 1.5 mL of isooctane/ethyl acetate 3:1 v/v 530 531 was added, vortexed vigorously, the organic phase was collected, and this step was repeated. The two volumes of organic phase were combined and taken to dryness by evaporation under 532 533 nitrogen gas at 40°C. Samples were resuspended in 300  $\mu$ L of isooctane/ethyl acetate 3:1 v/v. The diet samples were subsequently diluted 1:200 into isooctane/ethyl acetate 3:1 v/v. 534

535

#### 536 **Fatty acid quantification by GCMS**

Individual stable isotope fatty acid (FA) stock solutions were made in isooctane/ethyl acetate 3:1 537 v/v, a mixture containing 1.0  $\mu g/\mu L$  of every FA was made in isooctane/ethyl acetate 3:1 v/v that 538 539 was further diluted to 50 ng/ $\mu$ L, and stable isotope reference FA regression curves were prepared<sup>47.48</sup>. For total FA composition, 500 ng of the blended internal reference standard was 540 added to 50  $\mu$ L of total lipid extract, and samples were taken to dryness under N<sub>2</sub> gas. Dried 541 542 samples were immediately resuspended in 500  $\mu$ L of 100% ethanol, saponified with 500  $\mu$ L of 1 M NaOH at 90 °C for 45 min in Teflon capped tubes, and then acidified by addition of 525 µL of 1 543 544 M HCl. Saponified FA were re-extracted using 1 mL of isooctane (twice), dried under N<sub>2</sub> gas, and 545 were derivatized as above. The pentafluorobenzyl FA esters were resuspended in 200 µL of isooctane and diluted 1:10 into isooctane into GC/MS autosampler vials for injection. Analyte 546 547 data were acquired in NICI full scan, the FA-analyte peak area ratio to that of its corresponding 548 stable isotope reference FA was calculated for each analyte, and ratios were converted to absolute amounts relative to regression curves for each chain length and saturation<sup>48,49</sup>. Total 549 550 SFA, MUFA and PUFA was the quantitative sum of the nmoles of the class of fatty acid measured. Quantitative FA data were normalized to the total mass of diet input to the lipid extraction (i.e. 551 552 mg FA / g diet). Dietary FA amounts are listed in Table S4. Dietary FA intake (in mg) was calculated 553 by multiplying dietary FA amounts (mg/g) by the amount of diet consumed (g).

554

555 Data Availability

With respect to the IAEA database and the meta-analysis of BMR data this work comprises a secondary
analysis of data that are mostly already published and available in the primary literature. These data
have been compiled into a database, access to which is free. Forms for requesting data can be found at
www.dlwdatabase.org and should be directed to the lead corresponding author
<u>i.speakman@abdn.ac.uk</u> or Dr Alexia Alford at the (a.alford@iaea.org). The BMR data are available
upon request to co-corresponding author Dr Anura Kurpad (a.kurpad@sjri.res.in). The mouse data

described in the paper are available upon request to co-corresponding author Dr Matthew Rodeheffer(matthew.rodeheffer@yale.edu).

564

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# 580 Author contributions

JRS, KRW and LH processed and analysed the IAEA data, JMAdJ, JLK, and MCR collected, processed and analysed the mouse data, SS, SG, JRS and AK collected and analysed the retrospective BMR data from the literature. JRS, YY, HS, PNA, LFA, LJA, LA, IB, KBA, EEB, SB, AGB, CVCB, PB, MSB, NFB, SGJAC, GLC, JAC, RC, SKD, LRD, UE, SE, TF, BWF, AHG, MG, CH, AEH, MBH, SH, NJ, AMJ, PK, KPK, MK, WEK, RFK, EVL, AML, WRL, NL, CKM, ACM, EPM, JCM, JPM, MLN, TAN, RMO, HP, KHP, YPP, JPR, GP, RLP, RAR, SBR, DAR, ER, LMR, RMR, JR, SBR, MR, DAS, AJS, AMS, ES, SSU, GV, LMvE, EAvM, JCKW, GW, BMW, WWW, JAY, TY, XYZ contributed data to the database. JRS, YY, HS, SS, AJMM, CU, AHL, HP, JR, DAS and WWW created, curated and administered the database.

# **Conflict of interest**

591 The authors have no conflicts of interest to declare.

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**Table one**: Patterns of change in components of energy expenditure in males and females since the early 1990s. Data are shown unadjusted and adjusted for body composition and age. The gradient of the least squared regression fitted relationships with time are translated to the overall change in expenditure (MJ) over 30 years with the 95% confidence intervals (95%CI) for this change. TEE = total energy expenditure, BEE = basal energy expenditure, AEE = activity energy expenditure (=0.9TEE-BEE). Significance of the relationships is also shown. p > .05 was considered not significant (ns). All tests were two-sided.

# 619 **Males**

620	Unadjusted data			
621	Variable	Mean change over 30 y	95% CI	Significance
622		(MJ/d)	(± MJ/d)	
623	TEE	+0.55	0.73	0.138 (ns)
624	BEE	-1.19	0.536	p < .00002
625	AEE	+0.50	0.695	0.159 (ns)
626	Adjusted data			
627	TEE	-0.93	0.46	p < .0001
628	BEE	-0.96	0.15	p < 10 <sup>-9</sup>
629	AEE	+1.01	0.53	p < .0003
630				
631	Females			

632	Unadjusted data			
633 634	Variable	Mean change over 30 y (MJ/d)	95% CI	Significance
635	TEE	-0.16	0.360	0.405 (ns)
636	BEE	-0.32	0.352	0.071 (ns)
637	AEE	+0.18	0.452	0.448 (ns)

638 Adjusted data

639	TEE	-0.51	0.26	p < .00002
640	BEE	-0.12	0.215	0.276 (ns)
641	AEE	+0.42	0.367	p = 0.026

642

# 643 Figure legends

**Figure 1:** Trends over time for males in a) adjusted total energy expenditure, b) adjusted basal energy expenditure, and c) adjusted activity energy expenditure. Adjustments were made for body composition (fat and fat-free mass or body mass, and age) – see methods for details. All expenditures are in MJ/d and time is expressed in months since January 1982. Significant years are also indicated. Each data point is a different individual measurement of expenditure. The red lines are the fitted least squares regression fits. For regression details refer to text and Table 1.

Figure 2: Trends over time for females in a) adjusted total energy expenditure, b) adjusted basal
energy expenditure, and c) adjusted activity energy expenditure. Adjustments were made for
body composition (fat and lean mass and age) – see methods for details. Significant years are
also indicated. All expenditures are in MJ/d and time is expressed in months since January
1982. Each data point is a different individual measurement of expenditure. The red lines are
the fitted least squares regression fits. For regression details refer to text and Table 1.

Figure 3: A: effect of log<sub>e</sub> body mass on the log<sub>e</sub> basal metabolic rate (BMR) in a systematic review of 165 studies dating back to the early 1900s (first study 1919). Data for males in blue and for females in red. Studies with mixed male and female data not illustrated. B: Bubble plot showing the Residual log<sub>e</sub> Basal metabolism derived from a weighted regression of log<sub>e</sub> BMR against sex, age and log<sub>e</sub> (body mass) plotted against date of measurement in the same 165 studies. Bubbles represent the sample size of the studies. The red line is the fitted weighted regression.

- Figure 4: A: the relationship between body weight and metabolic rate in the mice fed different
  diets with variable fatty acid compositions. B: the effect of saturated fatty acid intake on
  residual metabolic rate corrected for body weight.
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